

A Modelled Cross-Correlation Search for Gravitational Waves from Scorpius X-1

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Abstract

The low-mass X-ray binary (LMXB) Scorpius X-1 (Sco X-1) is a promising source of gravitational waves in the advanced detector era. A variety of methods have been used or proposed to perform the directed search for gravitational waves from a binary source in a known sky location with unknown frequency and residual uncertainty in binary orbital parameters. One of these is modification of the cross-correlation method used in the stochastic search which takes into account the signal model of a rotating neutron star to allow cross-correlation of data from different times. By varying the maximum allowed time lag between cross-correlated segments, one can tune this semicoherent search and strike a balance between sensitivity and computing cost.

Gravitational Waves from LMXBs



Figure 1: Artist's impression of a low-mass X-ray binary. From *Astronomical Illustrations and Space Art*, by Fahad Sulehria, <http://www.novacelestia.com/>

A low-mass X-ray binary is a binary of a compact object (neutron star or black hole) & a companion star. If the CO is a NS, accretion from the companion can produce a hot spot & power GW emission from the non-axisymmetric NS. If GW spin-down balances accretion spinup, GW strength can be estimated from X-ray flux, and GW freq \approx constant [1]. Sco X-1, the brightest LMXB, is thought to be a $1.4M_{\odot}$ NS + $0.42M_{\odot}$ companion[2]. Proposed & applied search methods include a fully coherent search over a small amount of data [3], an unmodelled search for a monochromatic stochastic signal [4], a search for a pattern of sidebands arising from the Doppler modulation of the signal by the binary orbit [5], and the modelled cross-correlation search described here[6].

Cross-Correlation Method

- Divide data into segments of length T_{sft} & take “short Fourier transform” (SFT) $\tilde{x}_I(f)$.
- Label segments by I, J, \dots (I & J can be same or different times or detectors)
- Label pairs by α, β, \dots
- Use CW signal model ($\mathcal{A}_+ = \frac{1+\cos^2\iota}{2}; \mathcal{A}_\times = \cos\iota$)
$$h(t) = h_0 [\mathcal{A}_+ \cos\Phi(\tau(t))F_+ + \mathcal{A}_\times \sin\Phi(\tau(t))F_\times]$$
- expected cross-correlation btwn SFTs I & J
$$E[\tilde{x}_I^*(f_{k_I}) \tilde{x}_J(f_{k_J})] = \tilde{h}_I^*(f_{k_I}) \tilde{h}_J(f_{k_J}) = h_0^2 \tilde{\mathcal{G}}_{IJ} \delta_{T_{\text{sft}}}(f_{k_I} - f_I) \delta_{T_{\text{sft}}}(f_{k_J} - f_J)$$
- f_I is signal freq @ time T_I
Doppler shifted for detector I
- $\delta_{T_{\text{sft}}}(f - f') = \int_{-T_{\text{sft}}/2}^{T_{\text{sft}}/2} e^{i2\pi(f-f')t} dt$ so $\delta_{T_{\text{sft}}}(0) = T_{\text{sft}}$.
- Construct $\mathcal{Y}_{IJ} = \frac{\tilde{x}_I^*(f_{k_I}) \tilde{x}_J(f_{k_J})}{(T_{\text{sft}})^2}$ (where $f_{k_I} \approx f_I$) s.t.
$$E[\mathcal{Y}_{IJ}] \approx h_0^2 \tilde{\mathcal{G}}_{IJ} \quad \text{Var}[\mathcal{Y}_{IJ}] \approx \sigma_{IJ}^2 = \frac{S_I(f_0)S_J(f_0)}{4(T_{\text{sft}})^2}$$

- Optimally combine into $\rho = \frac{\sum_{\alpha} (u_{\alpha} \mathcal{Y}_{\alpha} + u_{\alpha}^* \mathcal{Y}_{\alpha}^*)}{w/u_{\alpha} \propto \frac{\tilde{\mathcal{G}}_{\alpha}^*}{\sigma_{\alpha}^2}}$ so $E[\rho] = h_0^2 \sqrt{2 \sum_{\alpha} \frac{|\tilde{\mathcal{G}}_{\alpha}|^2}{\sigma_{\alpha}^2}}$ & $\text{Var}[\rho] = 1$
- Computational considerations limit coherent integration time. Can make tunable semi-coherent search by restricting which SFT pairs α are included in $\rho = \sum_{\alpha} (u_{\alpha} \mathcal{Y}_{\alpha} + u_{\alpha}^* \mathcal{Y}_{\alpha}^*)$. E.g., only include pairs where $|T_I - T_J| \equiv |T_{\alpha}| \leq T_{\text{max}}$

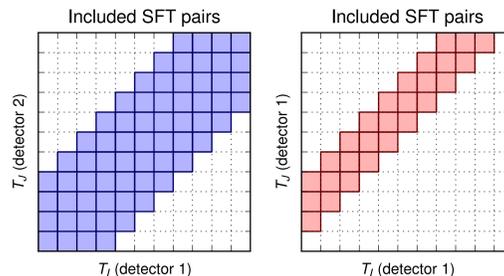


Figure 2: SFT pairs for inclusion in sliding cross-correlation search. Left: data from different detectors at same or different times. Right: data from same detector at different times. In this illustrative example, $T_{\text{max}} = 3T_{\text{sft}}$.

Parameter Space Metric

Consider dependence of ρ on parameters $\lambda \equiv \{\lambda_i\}$. Can define Parameter space metric via

$$\frac{E[\rho] - E[\rho^{\text{true}}]}{E[\rho^{\text{true}}]} = -g_{ij}(\Delta\lambda^i)(\Delta\lambda^j) + \mathcal{O}([\Delta\lambda]^3)$$

$$g_{ij} = -\frac{1}{2} \frac{E[\rho_{,ij}]|_{\lambda=\lambda^{\text{true}}}}{E[\rho^{\text{true}}]}$$

Assume dominant contribution to $E[\rho_{,ij}]$ is from variation of $\Delta\Phi_{IJ} = \Phi_I - \Phi_J$; get phase metric

$$g_{ij} = \frac{1}{2} \frac{\sum_{\alpha} \Delta\Phi_{\alpha,i} \Delta\Phi_{\alpha,j} |\tilde{\mathcal{G}}_{\alpha}|^2 / \sigma_{\alpha}^2}{\sum_{\beta} |\tilde{\mathcal{G}}_{\beta}|^2 / \sigma_{\beta}^2} \equiv \frac{1}{2} \langle \Delta\Phi_{\alpha,i} \Delta\Phi_{\alpha,j} \rangle_{\alpha}$$

Note $\langle \rangle_{\alpha}$ is average over pairs weighted by $\frac{|\tilde{\mathcal{G}}_{\alpha}|^2}{\sigma_{\alpha}^2}$; ignoring weighting factor gives usual metric [7]

$$\langle \Phi_{I,i} \Phi_{I,j} \rangle_I - \langle \Phi_{I,i} \rangle_I \langle \Phi_{I,j} \rangle_I$$

Define $T_{IJ} = T_I - T_J \equiv T_{\alpha}$ as time offset btwn SFTs; T_{α}^{av} is average time. For each detector pair, avg over pairs is avg over T_{α} & T_{α}^{av} . If we assume the avg over T_{α}^{av} evenly samples orbital phase, the metric in $\{f_0, a_p, \tilde{T}\}$ space is

$$\mathbf{g} = \begin{pmatrix} 2\pi^2 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{pmatrix} \langle T_{\alpha}^2 \rangle_{T_{\alpha}} + \begin{pmatrix} \pi^2 a_p^2 & \pi^2 f_0 a_p & 0 \\ \pi^2 f_0 a_p & \pi^2 f_0^2 & 0 \\ 0 & 0 & 4\pi^4 f_0^2 a_p^2 / P_{\text{orb}}^2 \end{pmatrix} \left\langle \sin^2 \frac{\pi T_{\alpha}}{P_{\text{orb}}} \right\rangle_{T_{\alpha}}$$

Since $\langle T_{\alpha}^2 \rangle_{T_{\alpha}} \gg a_p^2 \langle \sin^2 \frac{\pi T_{\alpha}}{P_{\text{orb}}} \rangle_{T_{\alpha}}$, metric approximately diagonal. Using this metric to place templates gives template count illustrated in Fig. 3.

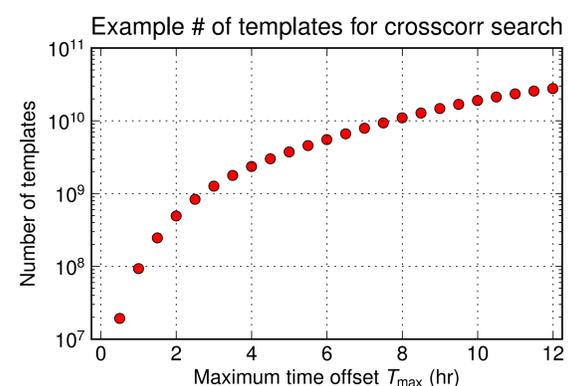


Figure 3: Illustration of dependence of template count on T_{max} ; *don't read too much into absolute numbers*

Sensitivity Estimates

Search is sensitive to signal of amplitude

$$h_0 = \left(\frac{S^2}{\sum_{\alpha} |\tilde{\mathcal{G}}_{\alpha}|^2 / \sigma_{\alpha}^2} \right)^2$$

where S is a statistical factor. $\tilde{\mathcal{G}}_{\alpha}$ depends on (unknown) spin orientation angles ι & ψ ; standard approach is to average value of $\tilde{\mathcal{G}}_{\alpha}$ over $\cos\iota$ & ψ . The ψ effect is small after average over sidereal time. The ι effect means actually

$$E[\rho] \approx h_0^2 \frac{A_+^2 + A_\times^2}{2} \sqrt{2 \sum_{\alpha} \frac{|\tilde{\mathcal{G}}_{\alpha}|^2}{\sigma_{\alpha}^2}}$$

Net effect is to change statistical factor S ; for 10% false-alarm & -dismissal, h_0 sensitivity is a factor of 1.4 worse. In Fig. 4 we illustrate the dependence of search sensitivity on the maximum allowed time offset T_{max}

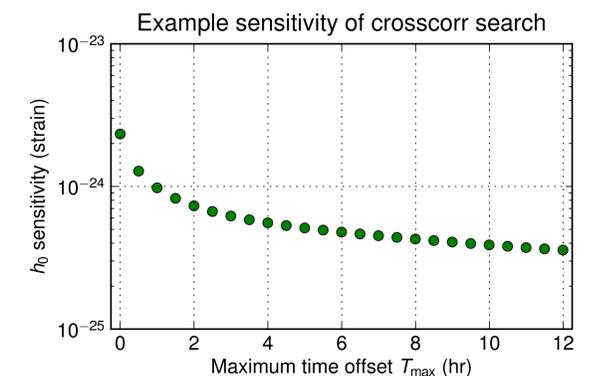


Figure 4: Illustration of dependence of template count on T_{max} ; *don't read too much into absolute numbers*. note that the $T_{\text{max}} = 0$ measurement is effectively the directed stochastic “radiometer” search.

As an illustration of the sensitivity of a practical search, we show in Fig. 5 the sensitivity with one year of data, using $T_{\text{max}} = 6$ hr.

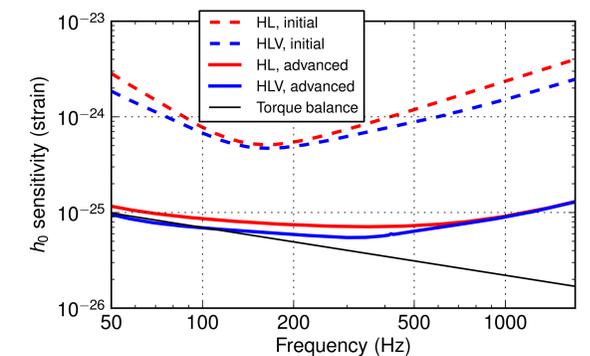


Figure 5: Sensitivity of a cross-correlation search with one year of LIGO and LIGO/Virgo design-sensitivity data, assuming 10% false-alarm & -dismissal, with $T_{\text{max}} = 6$ hr.

References

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